



Reference ranges for interrupter resistance technique: the Asthma UK Initiative

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ABSTRACT: Measuring interrupter resistance (R_{int}) is an increasingly popular lung function technique and especially suitable for preschool children because it is simple, quick and requires only passive cooperation. A European Respiratory Society (ERS)/American Thoracic Society (ATS) Task Force recently published empirical recommendations related to procedures, limitations and interpretation of the technique. However, for valid interpretation, high-quality reference equations are required and these have been lacking. The aim of the present study was to collate R_{int} data from healthy children in order to produce more robust reference equations. A further aim was to examine the influence of methodological differences on predicted R_{int} values.

R_{int} data from healthy children were collected from published and unpublished sources. Reference equations for expiratory and inspiratory R_{int} were developed using the LMS (lambda, mu, sigma) method.

Data from 1,090 children (51% males) aged 3–13 yrs were collated to construct sex-specific reference equations for expiratory R_{int} and data from 629 children (51% males) were collated for inspiratory R_{int} . Height was the best independent predictor of both expiratory and inspiratory R_{int} . Differences between centres were clinically irrelevant, and differences between ethnic groups could not be examined.

The availability of a large and generalisable sample and the use of modern statistical techniques enabled the development of more appropriate reference equations for R_{int} in young children.

KEYWORDS: Interrupter technique, preschool children, pulmonary function tests, reference equations

The measurement of airway resistance using the interrupter resistance (R_{int}) technique has become an increasingly popular lung function test in paediatric respiratory medicine. Since equipment for its measurement is commercially available and requires only passive cooperation, R_{int} provides a suitable lung function test for young children. The technique is safe, quick, noninvasive, available, inexpensive, applicable in field studies, and delivers results that are clinically relevant [1–12], appear sufficiently valid [13] and which seem suitable for assessing bronchodilator responses [7, 9, 14–16]. A European Respiratory Society (ERS)/American Thoracic Society (ATS) consensus statement was published in 2007, largely based on personal experience, in an attempt to make the procedure more uniform and facilitate comparisons between centres [17]. However,

interpretation of R_{int} remains limited and without a suitable reference population with which to compare results.

Reference equations are essential to express pulmonary function in relation to that which would be expected for healthy children of similar age, sex, body size and ethnic group. The choice of reference equation directly affects the interpretation of paediatric pulmonary function data, and this can have a major impact on patient care and research [18, 19]. So far, a clear recommendation regarding the best R_{int} reference equations has been lacking. Six published studies (reviewed in [20]) describing reference equations for R_{int} may have important limitations, because they are based on relatively small numbers of subjects (with <100 children <5 yrs of age), and it is unclear to what extent these can be generalised to other populations.

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Received:

Aug 05 2009

Accepted after revision:

Nov 27 2009

First published online:

Dec 23 2009

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It has previously been shown that it is feasible to collate existing reference data to produce generalisable reference equations for spirometry [21]. The present study aimed to collate available R_{int} reference data to produce more robust reference equations in young children. In addition, we examined the influence of methodological differences on predicted R_{int} values.

METHODS

The Asthma UK Collaborative Initiative (www.growinglungs.org.uk) was established to collate lung function data from healthy young children for spirometry, specific airway resistance and the interrupter technique. The present study focussed on expiratory R_{int} ($R_{int,exp}$) and inspiratory R_{int} ($R_{int,insp}$). Initially, the collaborative group was comprised of members of the ATS/ERS Task Force for paediatric pulmonary function testing. Subsequently, collaborators with access to pulmonary function data in healthy preschool children were identified by: systematically searching PubMed; advertising at international respiratory conferences; membership bulletins; word of mouth; and a manual search of relevant respiratory periodicals. R_{int} data were collected in healthy children aged 3–13 yrs, together with details regarding recruitment, population characteristics, equipment, measurement protocols and quality control. Collaborating centres were visited by one of the authors in order to conduct inter-lab comparisons, compare protocols and discuss methodological differences. All data were anonymised prior to contribution and came from research studies in which full local ethics approval and parental consent had been obtained.

Statistical methods for defining the reference range

The centile charts were constructed as described previously [21–23] using the LMS (lambda, mu, sigma) method [24]. This method is an extension of regression analysis, which includes three components: 1) the skewness (lambda), which models the departure of the variables from normality using a Box–Cox transformation; 2) the median (mu); and 3) the coefficient of variation (sigma), which models the spread of values around the median and adjusts for any non-uniform dispersion, hence LMS. The three quantities (LMS) are allowed to change with height and/or age, to reflect changes in the distribution as children grow. We applied the LMS method using the GAMLSS (Generalised Additive Models of Location, Shape and Scale) package in the statistical program R (version 2.6.1, R Foundation; available from www.r-project.org) [25].

During preliminary modelling, sex was independently associated with R_{int} and significant interactions were observed between sex and height; therefore, sex-specific models were created. According to previous reports, differences between $R_{int,insp}$ and $R_{int,exp}$ preclude combination of these data [3, 11], thereby necessitating separate reference equations. Fractional polynomials were used to fit each of the curves and explain the body size-related changes [26]. The goodness of fit was assessed using the Schwarz Bayesian Criterion, which compares consecutive models directly whilst adjusting for the increased complexity to determine the simplest model with best fit.

Prediction models

The fitted models provide sex-specific and height-adjusted values for the median, coefficient of variation and skewness. The median (M) is the predicted value for the individual

which, together with the coefficient of variation (S) and skewness (L), allows the individual's R_{int} measurement to be converted to a z-score:

$$z\text{-score} = ((\text{Measurement}/M)^{L-1}) / (L \times S)$$

$$\text{Upper limit of normal (ULN; 97.5th percentile)} = M \times (1.96 \times S \times L + 1)^{1/L}$$

z-scores are normally distributed with mean 0 and SD 1. Alternatively, the % predicted can simply be calculated from the median value, but this method of expressing the data is not recommended as it does not consider the between-subject variability of the measurements [17].

RESULTS

Eight centres with reference R_{int} data in healthy subjects were identified. Seven of these used the classical interruption technique with similar methodologies and produced similar results. The remaining study used a different measurement technique (the opening technique) [17] and procedure was noted as a clear outlier; this study was subsequently excluded. Data collection in one centre was still ongoing at the time of final data collation for the current initiative; it therefore could not be included. Results from five of the six centres who contributed R_{int} data have been published previously [3, 27–30]. In order to keep centres anonymous, a unique identifier was assigned to each centre and used thereafter. Data from healthy subjects (mean age 6.3 (range 2–13)) yrs, corresponding to a height range of 85–174 cm, with a similar proportion of males and females, were collated. One centre measured R_{int} in three ethnic groups but found no difference between the groups [29]. An additional centre measured R_{int} in a mixed-ethnic population but could not define ethnicity due to the highly diverse population characteristics [27]. All six centres contributed $R_{int,exp}$ measurements ($n=1,090$), with three centres also measuring $R_{int,insp}$ ($n=629$). In all centres, measurements were performed by the child breathing through a mouthpiece with a nose-clip *in situ*. A brief summary of the six studies is presented in table 1.

Population characteristics

All subjects included in the collated reference population were born at term and were free of chronic respiratory diseases, but the studies differed with respect to other exclusion criteria, such as history of wheeze (table 1). The majority of subjects had height-for-age, weight-for-age (where recorded) and body mass index-for-age z-scores between ± 2 SD [31].

Methodological differences

Three of the six centres were visited by one of the authors (P.J.F.M. Merkus). Methodologically, the three centres were similar. The remaining three centres completed a detailed questionnaire about the equipment and protocols.

$R_{int,exp}$ versus $R_{int,insp}$

In a subset of 535 subjects with both measurements, a comparison of $R_{int,insp}$ with $R_{int,exp}$ found no difference between the two (mean difference -0.001 (95% CI -0.016 – 0.013 , 95% limit of agreement -0.341 – 0.339) $\text{kPa}\cdot\text{L}^{-1}\cdot\text{s}^{-1}$). There was no relationship between the observed difference and age, height or weight; however, both the within-subject variability and absolute difference between these two outcomes tended to be slightly greater in younger children (*i.e.* those <110 cm) (fig. 1).

TABLE 1 General characteristics of the six studies (labelled A–F) included

Centre	Subjects n	Conditions	Equipment	Timing of shutter closure	Exclusion criteria
A	236	Lab $R_{int,exp}$	MicroLab 4000 and MicroRint	PETF; within 10 ms for 100 m	Acute disease or reported respiratory symptoms, chronic respiratory disease or symptoms, neurological diseases, respiratory tract infection within past 3 weeks
B	214	Lab $R_{int,insp}/R_{int,exp}$	MicroRint	PETF or PITF; within 10 ms for 100 ms	Respiratory symptoms 1 month prior to testing, eczema, allergy, parental smoking, doctor-diagnosed asthma, current asthma medication, cardiorespiratory disease or other chronic disease, anatomical abnormality, abnormality of the upper airway, vocal cord disorder
C	106	Field $R_{int,exp}$	MicroRint	PETF; within 10 ms for 100 ms	Wheeze in previous year, chronic respiratory disease, congenital heart disease, very low birth weight
D	91	Lab multicentre $R_{int,insp}/R_{int,exp}$	Spiroteq	Between 40 and 60% expiratory or inspiratory tidal volume; within 6 ms for 100 ms	Preterm birth (<36 weeks of gestation), intra-uterine growth restriction, chronic respiratory disease, cardiac disease, neurological disability, respiratory infection in past 4 weeks
E	193	Field $R_{int,exp}$	MicroRint	PETF; within 10 ms for 100 ms	Acute disease or reported respiratory symptoms, chronic respiratory disease, neurological or orthopaedic disability, preterm birth
F	284	Field $R_{int,insp}/R_{int,exp}$	MicroLab 4000	PETF or PITF; within 10 ms for 100 ms	History of wheeze ever, history of recent respiratory tract infection, preterm birth, chronic respiratory disease, cardiac disease, neurological disability

$R_{int,exp}$: expiratory interrupter resistance; $R_{int,insp}$: inspiratory interrupter resistance; PETF: peak expiratory tidal flow; PITF: peak inspiratory tidal flow. MicroLab 4000 and MicroRint are manufactured by MicroMedical® (Rochester, UK). Spiroteq is manufactured by Dyn'R® (Toulouse, France).

Trace exclusion

All contributors described a procedure to exclude incorrect measurements according to the child's behaviour, including: irregular breathing, vocal cord adduction and leakage around the mouthpiece. Five centres used additional criteria such as tachypnoea and extreme neck flexion/extension. All centres inspected individual mouth pressure–time recordings and discarded those deemed to be technically unacceptable. The criteria for such exclusions were not always clearly described,

with most investigators using an “eyeballing technique” to approve or discard pressure tracings [32].

Cheek support

All but one of the centres (F) supported the child's cheeks during R_{int} data collection. Centre F undertook a within-subject comparison with and without cheek support in 29 children for $R_{int,insp}$ and 39 children for $R_{int,exp}$ and did not find a significant difference between the two techniques. The 95% limits of agreement between the measurements with and without cheeks supported were -0.104 – 0.080 $\text{kPa}\cdot\text{L}^{-1}\cdot\text{s}^{-1}$ for $R_{int,insp}$ and -0.094 – 0.126 $\text{kPa}\cdot\text{L}^{-1}\cdot\text{s}^{-1}$ for $R_{int,exp}$ [28].

Filter use

Two centres (C and E) used the same bacterial filter. In both centres, the software took the filter resistance into account prior to reporting results.

Interruption trigger

Interruptions were triggered at peak tidal flow, in all but one of the centres (D), where interruptions occurred between 40–60% of tidal volume.

Method of R_{int} calculation

Mouth pressure was calculated using the two-points linear back extrapolation to 0 ms (centres A, B, C and E) or 15 ms (centres A, D and F) after the onset of interruption (table 1).

Number of tracings

The minimum number of technically acceptable tracings required per child (five to seven tracings) was similar in all

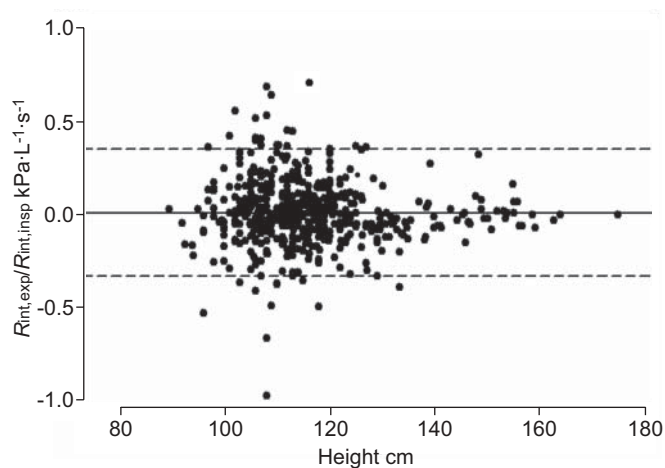


FIGURE 1. Bland–Altman plot of the difference between expiratory interrupter resistance ($R_{int,exp}$) and inspiratory interrupter resistance ($R_{int,insp}$) against height, demonstrating greater variability between measurements at younger ages.

centres. Data from children with fewer than five acceptable traces were excluded.

Mean or median

Four centres reported mean results, whereas two only reported the median. In the two centres in which $R_{int,exp}$ was reported both as mean and median (C and F, total 390 subjects), no difference was observed (mean 0.00 (95% CI -0.005–0.006, 95% limit of agreement -0.1–0.1) $kPa \cdot L^{-1} \cdot s^{-1}$). As these were interchangeable, data from centres that reported median results were combined with the other four centres.

Reference equations

The sex-specific reference equations for $R_{int,exp}$ and $R_{int,insp}$ are presented in table 2; $R_{int,exp}$ equations with 95% prediction limits are illustrated in figure 2. There were significant sex differences during preliminary modelling, with males having slightly lower $R_{int,exp}$ values; therefore, sex-specific equations were constructed. Using univariable analysis, height and age were both predictors of $R_{int,exp}$; using multivariable analysis, only height was independently associated with $R_{int,exp}$. Height was the only significant and independent predictor of $R_{int,insp}$. After adjustment for height, weight was not associated with either outcome and was not included in the final models. To avoid edge effects due to limited sample size at the extremes of the height range, we recommend that these equations are not extrapolated to children who are <90 cm or >160 cm in height.

The variability of measurements around the prediction equation was constant around the mean, and there was evidence of left skewness for both outcomes. Both of these observations have implications for defining the 97.5th percentile, the ULN. The estimates of the between-subject variability (S) and skewness (L) are presented in table 2; these, along with the median predicted value (M), can be combined algebraically

to calculate a z-score. For example, the $R_{int,exp}$ z-score and ULN for a male can be calculated using the following equation:

$$z\text{-score} = \frac{(\text{Observed } R_{int,exp})}{(0.1337 + 7631.3 * \text{Height}^{-2})^{0.4213} - 1} / (0.24 * 0.4213)$$

$$ULN = (0.1337 + (7631.3 * \text{Height}^{-2})) * (1.96 * 0.24 * 0.4213 + 1)^{1/0.4213}$$

$R_{int,exp}$ is expressed as $kPa \cdot L^{-1} \cdot s^{-1}$ and height is expressed as cm.

Between-centre differences

There was remarkable agreement between the centres and while on average some centres had z-scores that were above zero, and others where it was below zero, the magnitude of the offset was small (maximum 0.3 z-scores) and unlikely to affect clinical interpretation of results (fig. 3). The percentages of healthy children above the ULN in each centre were: A 2.5%; B 4.5%; C 3.3%; D 4.1%; E 4.7%; and F 0.1%.

DISCUSSION

The present study provides paediatric reference equations for $R_{int,exp}$ (n=1,090) and $R_{int,insp}$ (n=629) measurements from healthy children of 90–160 cm. The equations may facilitate more accurate interpretation of R_{int} results. By combining normative data from published and unpublished studies, the

TABLE 2		Reference equations for expiratory interrupter resistance ($R_{int,exp}$) and inspiratory interrupter resistance ($R_{int,insp}$)		
	M	S	L	
$R_{int,exp}$				
Males [#]	0.1337+7631.3*Height ⁻²	0.24	0.4213	
ULN	$(0.1337+(7631.3 * \text{Height}^{-2})) * (1.96 * 0.24 * 0.4213 + 1)^{1/0.4213}$			
Females [†]	0.1725+7281.0*Height ⁻²	0.25	0.3583	
ULN	$(0.1725+(7281.0 * \text{Height}^{-2})) * (1.96 * 0.25 * 0.3583 + 1)^{1/0.3583}$			
$R_{int,insp}$				
Males [‡]	0.0304+8586.0*Height ⁻²	0.25	0.4265	
ULN	$(0.0304+(8586.0 * \text{Height}^{-2})) * (1.96 * 0.25 * 0.4265 + 1)^{1/0.4265}$			
Females [§]	-0.01735+9304.0*Height ⁻²	0.24	0.2247	
ULN	$(-0.01735+(9304.0 * \text{Height}^{-2})) * (1.96 * 0.24 * 0.2247 + 1)^{1/0.2247}$			

M: mu; S: sigma; L: lambda; ULN: upper limit of normal. [#]: n=553; [†]: n=537; [‡]: n=321; [§]: n=308.

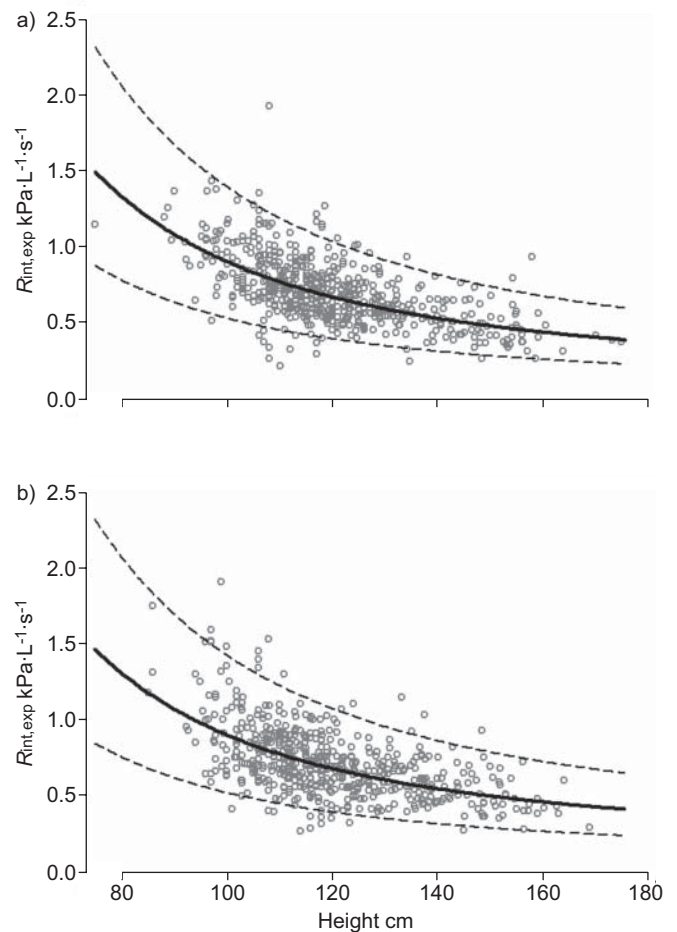


FIGURE 2. Sex-specific reference equations for expiratory interrupter resistance ($R_{int,exp}$; —) and 95% prediction limits (---) for a) males and b) females.

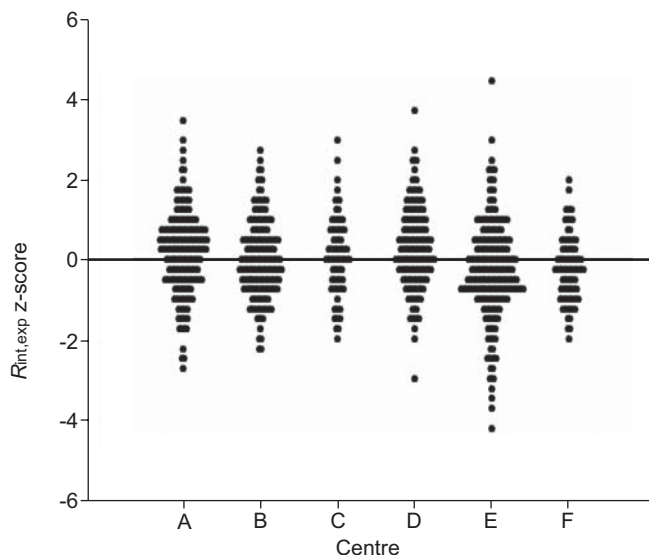


FIGURE 3. Centre-specific z-scores for expiratory interrupter resistance ($R_{int,exp}$) demonstrated good agreement between the six centres, despite some offset. Mean (95% CI) z-scores in each centre: A 0.15 (0.02–0.27); B -0.03 (-0.15–0.09); C -0.27 (-0.41–-0.11); D 0.05 (-0.15–0.24); E 0.25 (0.11–0.39); F -0.29 (-0.42–-0.16).

collated dataset describes a generalisable population, making the reference applicable across a variety of populations. In addition, the larger sample size allowed investigation of the distributional characteristics of the data, and exploration of various demographic variables. The use of the LMS method allowed for a more accurate definition of the ULN, which may improve the technique's discrimination between health and disease. Further work is needed to validate the ULN in different disease groups, particularly those for which R_{int} is most useful clinically.

These equations are appropriate when: 1) the classic interrupter technique is performed using commercially available equipment; 2) children are seated, wearing nose-clips and using a mouthpiece; 3) children have their cheeks supported; and 4) children are ≥ 90 cm and ≤ 160 cm tall.

For the first time, separate equations are presented for males and females. The larger sample allowed detection of statistically significant differences between the sexes; however, the observed differences were small (table 3) and unlikely to affect clinical interpretation of results. While we would expect to observe age-dependent sex differences, reflecting the differential development of the airways between females and males during growth [33], we found that females had higher resistance at all ages. While this was unexpected, it may reflect the fact that R_{int} probably measures central airway function, whereas the enhanced airway function reported in females prior to puberty is most noticeable in indices that reflect more peripheral airway calibre, such as forced expired flows [21, 34].

We present reference equations for both $R_{int,exp}$ and $R_{int,insp}$ as previous studies have shown that these two outcomes are not interchangeable [3, 11]. In a subset of 535 children, with both outcomes we found minimal differences between these two.

TABLE 3 Examples of predicted values and upper limits of normal (ULN) for expiratory interrupter resistance ($\text{kPa}\cdot\text{L}^{-1}\cdot\text{s}^{-1}$) in males and females of various heights

Height cm	Males		Females	
	Predicted	ULN	Predicted	ULN
90	1.09	1.67	1.07	1.68
100	0.89	1.37	0.90	1.41
110	0.76	1.17	0.77	1.21
120	0.66	1.01	0.68	1.07
130	0.59	0.91	0.60	0.94
140	0.52	0.80	0.54	0.85
150	0.47	0.73	0.51	0.80
160	0.43	0.66	0.46	0.74

The increased within-subject variability between $R_{int,insp}$ and $R_{int,exp}$ in younger children could either reflect higher airway compliance during tidal breathing and/or slightly less consistent measurement technique in this age group, as observed for spirometry [21].

The retrospective nature of the data collection meant there were differences in population characteristics (exclusion criteria) between the six centres. For instance, whereas all centres excluded preterm children, not all excluded children with a history of wheeze or passive smoke exposure. The collated population therefore lies somewhere between a reference and an ideal population, and reflects a typical testing population, which may be more generalisable across different populations. The dataset was also limited in that we could not investigate ethnic differences. It is likely that the current study, as well as available literature, is underpowered to detect such differences. A recently published study describing reference equations for a Chinese population noted reference values similar to other published studies in Caucasian children >110 cm; however, the differences observed in shorter children may reflect the limited number of subjects <110 cm [35].

The retrospective nature of the study also limited our ability to separate methodological differences from population differences. While methodologically the centres were similar with respect to equipment used and measurement techniques, subtle methodological differences may explain some of the observed differences in the measured values. In a subset of subjects, we were able to demonstrate that results summarised as mean were interchangeable with results presented as the median. A similar conclusion was reached when direct within-subject comparisons were investigated in a group of 40 healthy children [36]. Slightly lower predicted values and somewhat larger between-subject variability were found in centre F (fig. 2), where cheeks were not supported during the measurements. However, we consider it unlikely that cheek support played a relevant role in this study as all the children were healthy and had no signs of airway obstruction. As suggested in the recent ERS/ATS guidelines [17], we recommend that in future studies cheeks should be supported during measurement

as this decreases variability and increases accuracy of measurement in children with airway obstruction. Most centres used eyeballing to determine quality of the measurement [32]. Although one could argue that eyeballing as a quality control measure is somewhat subjective, a previous study found clinical interpretation of the measurement was not affected, and that fewer curves were rejected compared to using objective quality criteria [32].

It is important to note that the entire collated dataset of the present study was based on cross-sectional measurements, which may not be appropriate for interpreting longitudinal changes. Nonetheless, these reference equations should be valid for interpreting cross-sectional data, as long as measurements are similar to those used in the current study. If equipment or measurement techniques change, or other ethnic groups are investigated, these reference equations will need to be updated.

Conclusions

These collated reference equations provide a step towards improved interpretation of R_{int} measurements in young children. We recommend that manufacturers of R_{int} equipment incorporate these reference equations into their software, so that those using the same techniques, and conducting measurements in similar populations according to the same international guidelines [17], can optimise interpretation of results. These equations should also facilitate more meaningful comparisons between centres. This may create new opportunities for future multicentre studies in young children with respiratory disease, in which R_{int} is the selected outcome.

SUPPORT STATEMENT

This study was funded by Asthma UK. J. Kivastik was funded by the Estonian Science Foundation (grant no. 7322).

STATEMENT OF INTEREST

A statement of interest for S.A. McKenzie can be found at www.ersjournals.com/misc/statements.dtl

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